Linear Models for Regression Greg Mori - CMPT 419/726

Bishop PRML Ch. 3

Outline

Regression

Linear Basis Function Models

Loss Functions for Regression

Finding Optimal Weights

Regularization

Bayesian Linear Regression

Outline

Regression

Linear Basis Function Models

Loss Functions for Regression

Finding Optimal Weights

Regularization

Bayesian Linear Regression

Regression



- Given training set $\{(x_1, t_1), \ldots, (x_N, t_N)\}$
- t_i is continuous: regression
- For now, assume $t_i \in \mathbb{R}$, $x_i \in \mathbb{R}^D$
- E.g. t_i is stock price, x_i contains company profit, debt, cash flow, gross sales, number of spam emails sent, . . .

Outline

Regression

Linear Basis Function Models

Loss Functions for Regression

Finding Optimal Weights

Regularization

Bayesian Linear Regression

Linear Functions

• A function $f(\cdot)$ is linear if

$$f(\alpha u + \beta v) = \alpha f(u) + \beta f(v)$$

 Linear functions will lead to simple algorithms, so let's see what we can do with them

Linear Regression

• Simplest linear model for regression

$$y(x, w) = w_0 + w_1x_1 + w_2x_2 + \ldots + w_Dx_D$$

- Remember, we're learning w
- Set w so that y(x, w) aligns with target value in training data

Linear Regression

Simplest linear model for regression

$$y(x, w) = w_0 + w_1x_1 + w_2x_2 + \ldots + w_Dx_D$$

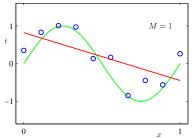
- Remember, we're learning w
- Set w so that y(x, w) aligns with target value in training data

Linear Regression

Simplest linear model for regression

$$y(x, w) = w_0 + w_1x_1 + w_2x_2 + \ldots + w_Dx_D$$

- Remember, we're learning w
- Set w so that y(x, w) aligns with target value in training data
- This is a very simple model, limited in what it can do



· Simplest linear model

$$y(x, w) = w_0 + w_1 x_1 + w_2 x_2 + \ldots + w_D x_D$$

was linear in x (*) and w

- Linear in w is what will be important for simple algorithms
- Extend to include fixed non-linear functions of data

$$y(x, w) = w_0 + w_1\phi_1(x) + w_2\phi_2(x) + \ldots + w_{M-1}\phi_{M-1}(x)$$

 Linear combinations of these basis functions also linear in parameters

Simplest linear model

$$y(x, w) = w_0 + w_1 x_1 + w_2 x_2 + \ldots + w_D x_D$$

was linear in x (*) and w

- Linear in w is what will be important for simple algorithms
- Extend to include fixed non-linear functions of data

$$y(x, w) = w_0 + w_1\phi_1(x) + w_2\phi_2(x) + \ldots + w_{M-1}\phi_{M-1}(x)$$

• Linear combinations of these basis functions also linear in parameters

Simplest linear model

$$y(\mathbf{x}, \mathbf{w}) = w_0 + w_1 x_1 + w_2 x_2 + \ldots + w_D x_D$$

was linear in x (*) and w

- Linear in w is what will be important for simple algorithms
- Extend to include fixed non-linear functions of data

$$y(\mathbf{x}, \mathbf{w}) = w_0 + w_1 \phi_1(\mathbf{x}) + w_2 \phi_2(\mathbf{x}) + \ldots + w_{M-1} \phi_{M-1}(\mathbf{x})$$

 Linear combinations of these basis functions also linear in parameters

Bias parameter allows fixed offset in data

$$y(\mathbf{x}, \mathbf{w}) = \underbrace{w_0}_{bias} + w_1 \phi_1(\mathbf{x}) + w_2 \phi_2(\mathbf{x}) + \dots + w_{M-1} \phi_{M-1}(\mathbf{x})$$

• Think of simple 1-D x:

$$y(x, w) = \underbrace{w_0}_{intercept} + \underbrace{w_1}_{slope} x$$

• For notational convenience, define $\phi_0(x) = 1$:

$$y(\mathbf{x}, \mathbf{w}) = \sum_{i=0}^{M-1} w_i \phi_i(\mathbf{x}) = \mathbf{w}^T \phi(\mathbf{x})$$

Bias parameter allows fixed offset in data

$$y(\mathbf{x}, \mathbf{w}) = \underbrace{w_0}_{bias} + w_1 \phi_1(\mathbf{x}) + w_2 \phi_2(\mathbf{x}) + \dots + w_{M-1} \phi_{M-1}(\mathbf{x})$$

• Think of simple 1-D x:

$$y(x, w) = \underbrace{w_0}_{intercept} + \underbrace{w_1}_{slope} x$$

• For notational convenience, define $\phi_0(x) = 1$:

$$y(\mathbf{x}, \mathbf{w}) = \sum_{i=0}^{M-1} w_i \phi_i(\mathbf{x}) = \mathbf{w}^T \phi(\mathbf{x})$$

· Bias parameter allows fixed offset in data

$$y(\mathbf{x}, \mathbf{w}) = \underbrace{w_0}_{bias} + w_1 \phi_1(\mathbf{x}) + w_2 \phi_2(\mathbf{x}) + \dots + w_{M-1} \phi_{M-1}(\mathbf{x})$$

• Think of simple 1-D x:

$$y(x, w) = \underbrace{w_0}_{intercept} + \underbrace{w_1}_{slope} x$$

• For notational convenience, define $\phi_0(x) = 1$:

$$y(\mathbf{x}, \mathbf{w}) = \sum_{j=0}^{M-1} w_j \phi_j(\mathbf{x}) = \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x})$$

• Function for regression y(x, w) is non-linear function of x, but linear in w:

$$y(\mathbf{x}, \mathbf{w}) = \sum_{j=0}^{M-1} w_j \phi_j(\mathbf{x}) = \mathbf{w}^T \phi(\mathbf{x})$$

- Polynomial regression is an example of this
- Order *M* polynomial regression, $\phi_i(x) = ?$

• Function for regression y(x, w) is non-linear function of x, but linear in w:

$$y(\mathbf{x}, \mathbf{w}) = \sum_{i=0}^{M-1} w_j \phi_j(\mathbf{x}) = \mathbf{w}^T \phi(\mathbf{x})$$

- · Polynomial regression is an example of this
- Order *M* polynomial regression, $\phi_i(x) = ?$

 Function for regression y(x, w) is non-linear function of x, but linear in w:

$$y(\mathbf{x}, \mathbf{w}) = \sum_{j=0}^{M-1} w_j \phi_j(\mathbf{x}) = \mathbf{w}^T \phi(\mathbf{x})$$

- Polynomial regression is an example of this
- Order *M* polynomial regression, $\phi_i(x) = ?$
- $\phi_i(x) = x^j$:

$$y(x, w) = w_0 x^0 + w_1 x^1 + \ldots + w_M x^M$$

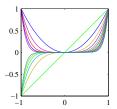
- Often we extract features from x
 - An intuitve way to think of $\phi_i(x)$ is as feature functions
- E.g. Automatic CMPT726 project report grading system
 - x is text of report: In this project we apply the algorithm of Mori [2] to recognizing blue objects. We test this algorithm on pictures of you and I from my holiday photo collection. ...
- $\phi_1(x)$ is count of occurrences of Mori [
- $\phi_2(x)$ is count of occurrences of of you and I
- Regression grade $y(x, w) = 20\phi_1(x) 10\phi_2(x)$

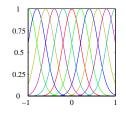
- Often we extract features from x
 - An intuitve way to think of $\phi_i(x)$ is as feature functions
- E.g. Automatic CMPT726 project report grading system
 - x is text of report: In this project we apply the algorithm of Mori [2] to recognizing blue objects. We test this algorithm on pictures of you and I from my holiday photo collection. ...
- $\phi_1(x)$ is count of occurrences of Mori [
- $\phi_2(x)$ is count of occurrences of of you and I
- Regression grade $y(x, w) = 20\phi_1(x) 10\phi_2(x)$

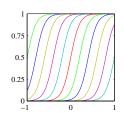
- Often we extract features from x
 - An intuitve way to think of $\phi_i(x)$ is as feature functions
- E.g. Automatic CMPT726 project report grading system
 - x is text of report: In this project we apply the algorithm of Mori [2] to recognizing blue objects. We test this algorithm on pictures of you and I from my holiday photo collection. ...
- $\phi_1(x)$ is count of occurrences of Mori [
- $\phi_2(x)$ is count of occurrences of of you and I
- Regression grade $y(x, w) = 20\phi_1(x) 10\phi_2(x)$

- Often we extract features from x
 - An intuitve way to think of $\phi_i(x)$ is as feature functions
- E.g. Automatic CMPT726 project report grading system
 - x is text of report: In this project we apply the algorithm of Mori [2] to recognizing blue objects. We test this algorithm on pictures of you and I from my holiday photo collection. ...
- $\phi_1(x)$ is count of occurrences of Mori [
- $\phi_2(x)$ is count of occurrences of of you and I
- Regression grade $y(x, w) = 20\phi_1(x) 10\phi_2(x)$

Other Non-linear Basis Functions







- Polynomial $\phi_i(x) = x^j$
- Gaussians $\phi_j(x) = \exp\{-\frac{(x-\mu_j)^2}{2s^2}\}$
- Sigmoidal $\phi_j(x) = \frac{1}{1 + \exp((\mu_i x)/s)}$



• Use Gaussian basis functions, regression on temperature



• μ_1 = Vancouver, μ_2 = San Francisco, μ_3 = Oakland



- μ_1 = Vancouver, μ_2 = San Francisco, μ_3 = Oakland
- Temperature in x = Seattle? $y(x, w) = w_1 \exp\{-\frac{(x-\mu_1)^2}{2s^2}\} + w_2 \exp\{-\frac{(x-\mu_2)^2}{2s^2}\} + w_3 \exp\{-\frac{(x-\mu_3)^2}{2s^2}\}$



- μ_1 = Vancouver, μ_2 = San Francisco, μ_3 = Oakland
- Temperature in x = Seattle? $y(x, w) = w_1 \exp\{-\frac{(x-\mu_1)^2}{2s^2}\} + w_2 \exp\{-\frac{(x-\mu_2)^2}{2s^2}\} + w_3 \exp\{-\frac{(x-\mu_3)^2}{2s^2}\}$
- Compute distances to all μ , $y(x, w) \approx w_1$



Example - Gaussian Basis Functions: 726 Report Grading

Define:

- μ_1 = Crime and Punishment
- $\mu_2 = \text{Animal Farm}$
- $\mu_3 =$ Some paper by Mori
- Learn weights:
 - $w_1 = ?$
 - $w_2 = ?$
 - $w_3 = ?$
- Grade a project report x:
 - Measure similarity of x to each μ , Gaussian, with weights: $y(x, w) = \frac{(x_1, y_1)^2}{(x_1, y_2)^2}$
- The Gaussian basis function models end up similar to template matching

Example - Gaussian Basis Functions: 726 Report Grading

Define:

- μ_1 = Crime and Punishment
- $\mu_2 = \text{Animal Farm}$
- $\mu_3 =$ Some paper by Mori

· Learn weights:

- $w_1 = ?$
- $w_2 = ?$
- $w_3 = ?$
- Grade a project report x:
 - Measure similarity of x to each μ , Gaussian, with weights: $y(x, w) = w_1 \exp\{-\frac{(x-\mu_1)^2}{2}\} + w_2 \exp\{-\frac{(x-\mu_2)^2}{2}\} + w_3 \exp\{-\frac{(x-\mu_3)^2}{2}\}$
- The Gaussian basis function models end up similar to template matching

Example - Gaussian Basis Functions: 726 Report Grading

- Define:
 - μ_1 = Crime and Punishment
 - $\mu_2 = Animal Farm$
 - μ_3 = Some paper by Mori
- · Learn weights:
 - $w_1 = ?$
 - $w_2 = ?$
 - $w_3 = ?$
- Grade a project report x:
 - Measure similarity of x to each μ , Gaussian, with weights:

$$y(\mathbf{x}, \mathbf{w}) = w_1 \exp\{-\frac{(x-\mu_1)^2}{2s^2}\} + w_2 \exp\{-\frac{(x-\mu_2)^2}{2s^2}\} + w_3 \exp\{-\frac{(x-\mu_3)^2}{2s^2}\}$$

 The Gaussian basis function models end up similar to template matching

Outline

Regression

Linear Basis Function Models

Loss Functions for Regression

Finding Optimal Weights

Regularization

Bayesian Linear Regression

Loss Functions for Regression

- We want to find the "best" set of coefficients w
- Recall, one way to define "best" was minimizing squared error:

$$E(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} \{ y(x_n, \mathbf{w}) - t_n \}^2$$

We will now look at another way, based on maximum likelihood

Gaussian Noise Model for Regression

- We are provided with a training set $\{(x_i, t_i)\}$
- Let's assume t arises from a deterministic function plus Gassian distributed (with precision β) noise:

$$t = y(\mathbf{x}, \mathbf{w}) + \epsilon$$

• The probability of observing a target value *t* is then:

$$p(t|\mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t|\mathbf{y}(\mathbf{x}, \mathbf{w}), \beta^{-1})$$

• Notation: $\mathcal{N}(x|\mu, \sigma^2)$; x drawn from Gaussian with mean μ , variance σ^2

Gaussian Noise Model for Regression

- We are provided with a training set $\{(x_i, t_i)\}$
- Let's assume t arises from a deterministic function plus Gassian distributed (with precision β) noise:

$$t = y(\mathbf{x}, \mathbf{w}) + \epsilon$$

The probability of observing a target value t is then:

$$p(t|\mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t|y(\mathbf{x}, \mathbf{w}), \beta^{-1})$$

• Notation: $\mathcal{N}(x|\mu,\sigma^2)$; x drawn from Gaussian with mean μ , variance σ^2

Maximum Likelihood for Regression

• The likelihood of data $t = \{t_i\}$ using this Gaussian noise model is:

$$p(t|\mathbf{w},\beta) = \prod_{n=1}^{N} \mathcal{N}(t_n|\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1})$$

The log-likelihood is:

$$\ln p(t|w,\beta) = \ln \prod_{n=1}^{N} \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp(-\frac{\beta}{2} (t_n - w^T \phi(x_n))^2)$$

$$= \underbrace{\frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi)}_{const. wrt w} - \beta \underbrace{\frac{1}{2} \sum_{n=1}^{N} (t_n - w^T \phi(x_n))^2}_{const. wrt w}$$

 Sum of squared errors is maximum likelihood under a Gaussian noise model



Maximum Likelihood for Regression

• The likelihood of data $t = \{t_i\}$ using this Gaussian noise model is:

$$p(t|\mathbf{w},\beta) = \prod_{n=1}^{N} \mathcal{N}(t_n|\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1})$$

The log-likelihood is:

$$\ln p(t|\mathbf{w},\beta) = \ln \prod_{n=1}^{N} \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp(-\frac{\beta}{2} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2)$$

$$= \underbrace{\frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi)}_{const. \ wrt \ w} - \beta \underbrace{\frac{1}{2} \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2}_{squared \ error}$$

 Sum of squared errors is maximum likelihood under a Gaussian noise model



Maximum Likelihood for Regression

• The likelihood of data $t = \{t_i\}$ using this Gaussian noise model is:

$$p(t|\mathbf{w},\beta) = \prod_{n=1}^{N} \mathcal{N}(t_n|\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1})$$

The log-likelihood is:

$$\ln p(t|\mathbf{w},\beta) = \ln \prod_{n=1}^{N} \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp(-\frac{\beta}{2} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2)$$

$$= \underbrace{\frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi)}_{const. \ wrt \ \mathbf{w}} - \beta \underbrace{\frac{1}{2} \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2}_{squared \ error}$$

 Sum of squared errors is maximum likelihood under a Gaussian noise model



Maximum Likelihood for Regression

• The likelihood of data $t = \{t_i\}$ using this Gaussian noise model is:

$$p(t|\mathbf{w},\beta) = \prod_{n=1}^{N} \mathcal{N}(t_n|\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1})$$

The log-likelihood is:

$$\ln p(t|\mathbf{w},\beta) = \ln \prod_{n=1}^{N} \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp(-\frac{\beta}{2} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2)$$

$$= \underbrace{\frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi)}_{const. \ wrt \ \mathbf{w}} - \beta \underbrace{\frac{1}{2} \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2}_{squared \ error}$$

 Sum of squared errors is maximum likelihood under a Gaussian noise model



Outline

Regression

Linear Basis Function Models

Loss Functions for Regression

Finding Optimal Weights

Regularization

- How do we maximize likelihood wrt w (or minimize squared error)?
- Take gradient of log-likelihood wrt w:

$$\frac{\partial}{\partial w_i} \ln p(t|\mathbf{w}, \beta) = \beta \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n)) \phi_i(\mathbf{x}_n)$$

In vector form:

$$\nabla \ln p(t|\mathbf{w},\beta) = \beta \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n)) \phi(\mathbf{x}_n)^T$$

- How do we maximize likelihood wrt w (or minimize squared error)?
- Take gradient of log-likelihood wrt w:

$$\frac{\partial}{\partial w_i} \ln p(t|\mathbf{w}, \beta) = \beta \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n)) \phi_i(\mathbf{x}_n)$$

In vector form:

$$\nabla \ln p(t|\mathbf{w},\beta) = \beta \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n)) \phi(\mathbf{x}_n)^T$$

- How do we maximize likelihood wrt w (or minimize squared error)?
- Take gradient of log-likelihood wrt w:

$$\frac{\partial}{\partial w_i} \ln p(t|\mathbf{w}, \beta) = \beta \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n)) \phi_i(\mathbf{x}_n)$$

In vector form:

$$\nabla \ln p(t|\mathbf{w},\beta) = \beta \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n)) \phi(\mathbf{x}_n)^T$$

• Set gradient to 0:

$$\mathbf{0}^T = \sum_{n=1}^N t_n \phi(\mathbf{x}_n)^T - \mathbf{w}^T \left(\sum_{n=1}^N \phi(\mathbf{x}_n) \phi(\mathbf{x}_n)^T \right)$$

Maximum likelihood estimate for w:

$$egin{aligned} oldsymbol{w}_{ML} &= \left(oldsymbol{\Phi}^Toldsymbol{\Phi}
ight)^{-1}oldsymbol{\Phi}^T t \ oldsymbol{\Phi} &= \left(egin{array}{cccc} \phi_0(x_1) & \phi_1(x_1) & \dots & \phi_{M-1}(x_1) \ \phi_0(x_2) & \phi_1(x_2) & \dots & \phi_{M-1}(x_2) \ dots & dots & \ddots & dots \ \phi_0(x_N) & \phi_1(x_N) & \dots & \phi_{M-1}(x_N) \end{array}
ight) \end{aligned}$$

• $\Phi^{\dagger} = (\Phi^T \Phi)^{-1} \Phi^T$ known as the pseudo-inverse (numpy.linalg.pinv in python)

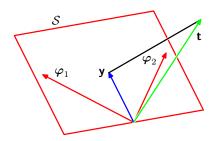
Set gradient to 0:

$$\mathbf{0}^T = \sum_{n=1}^N t_n \boldsymbol{\phi}(\mathbf{x}_n)^T - \mathbf{w}^T \left(\sum_{n=1}^N \boldsymbol{\phi}(\mathbf{x}_n) \boldsymbol{\phi}(\mathbf{x}_n)^T \right)$$

Maximum likelihood estimate for w:

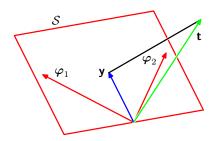
$$egin{aligned} oldsymbol{w}_{ML} &= \left(oldsymbol{\Phi}^Toldsymbol{\Phi}^Toldsymbol{t} \ oldsymbol{\Phi} &= \left(egin{array}{cccc} \phi_0(oldsymbol{x}_1) & \phi_1(oldsymbol{x}_1) & \dots & \phi_{M-1}(oldsymbol{x}_1) \ \phi_0(oldsymbol{x}_2) & \phi_1(oldsymbol{x}_2) & \dots & \phi_{M-1}(oldsymbol{x}_2) \ dots & dots & \ddots & dots \ \phi_0(oldsymbol{x}_N) & \phi_1(oldsymbol{x}_N) & \dots & \phi_{M-1}(oldsymbol{x}_N) \end{array}
ight) \end{aligned}$$

• $\Phi^{\dagger} = (\Phi^T \Phi)^{-1} \Phi^T$ known as the pseudo-inverse (numpy.linalg.pinv in python)



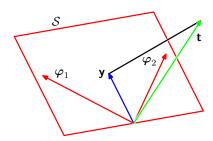
- $t = (t_1, \dots, t_N)$ is the target value vector
- S is space spanned by $\varphi_i = (\phi_i(x_1), \dots, \phi_i(x_N))$
- Solution y lies in S
- Least squares solution is orthogonal projection of t onto S
- Can verify this by looking at $y = \Phi w_{ML} = \Phi \Phi^\dagger t = Pt$

•
$$P^2 = P \cdot P = P^T$$



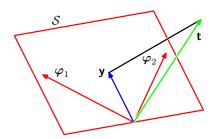
- $t = (t_1, \dots, t_N)$ is the target value vector
- \mathcal{S} is space spanned by $\boldsymbol{\varphi}_j = (\phi_j(\boldsymbol{x}_1), \dots, \phi_j(\boldsymbol{x}_N))$
- Solution y lies in S
- Least squares solution is orthogonal projection of t onto S
- Can verify this by looking at $y = \Phi w_{ML} = \Phi \Phi^\dagger t = Pt$

•
$$P^2 = P$$
. $P = P^T$



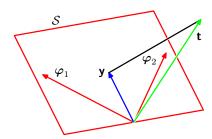
- $t = (t_1, \dots, t_N)$ is the target value vector
- S is space spanned by $\varphi_i = (\phi_j(x_1), \dots, \phi_j(x_N))$
- Solution y lies in S
- Least squares solution is orthogonal projection of t onto S
- Can verify this by looking at $y = \Phi w_{ML} = \Phi \Phi^\dagger t = Pt$

•
$$P^2 = P$$
. $P = P^T$



- $t = (t_1, \dots, t_N)$ is the target value vector
- S is space spanned by $\varphi_j = (\phi_j(x_1), \dots, \phi_j(x_N))$
- Solution y lies in S
- Least squares solution is orthogonal projection of t onto S
- Can verify this by looking at $y = \Phi w_{ML} = \Phi \Phi^\dagger t = Pt$

•
$$P^2 = P$$
. $P = P^T$



- $t = (t_1, \dots, t_N)$ is the target value vector
- S is space spanned by $\varphi_j = (\phi_j(x_1), \dots, \phi_j(x_N))$
- Solution y lies in S
- Least squares solution is orthogonal projection of t onto ${\cal S}$
- Can verify this by looking at $y = \Phi w_{ML} = \Phi \Phi^\dagger t = Pt$

•
$$P^2 = P$$
, $P = P^T$

- In practice N might be huge, or data might arrive online
- Can use a gradient descent method:
 - Start with initial guess for w
 - Update by taking a step in gradient direction ∇E of error function
- Modify to use stochastic / sequential gradient descent:
 - If error function $E = \sum_n E_n$ (e.g. least squares)
 - Update by taking a step in gradient direction ∇E_n for one example
 - Details about step size are important decrease step size at the end

- In practice N might be huge, or data might arrive online
- · Can use a gradient descent method:
 - Start with initial guess for w
 - Update by taking a step in gradient direction ∇E of error function
- Modify to use stochastic / sequential gradient descent:
 - If error function $E = \sum_n E_n$ (e.g. least squares)
 - Update by taking a step in gradient direction ∇E_n for one example
 - Details about step size are important decrease step size at the end

- In practice N might be huge, or data might arrive online
- · Can use a gradient descent method:
 - Start with initial guess for w
 - Update by taking a step in gradient direction ∇E of error function
- Modify to use stochastic / sequential gradient descent:
 - If error function $E = \sum_n E_n$ (e.g. least squares)
 - Update by taking a step in gradient direction ∇E_n for one example
 - Details about step size are important decrease step size at the end

- In practice N might be huge, or data might arrive online
- · Can use a gradient descent method:
 - Start with initial guess for w
 - Update by taking a step in gradient direction ∇E of error function
- Modify to use stochastic / sequential gradient descent:
 - If error function $E = \sum_n E_n$ (e.g. least squares)
 - Update by taking a step in gradient direction ∇E_n for one example
 - Details about step size are important decrease step size at the end

Outline

Regression

Linear Basis Function Models

Loss Functions for Regression

Finding Optimal Weights

Regularization

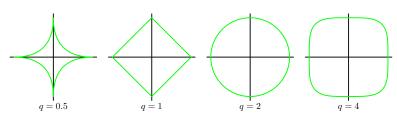
Regularization

 Last week we discussed regularization as a technique to avoid over-fitting:

$$\tilde{E}(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} \{ y(x_n, \mathbf{w}) - t_n \}^2 + \underbrace{\frac{\lambda}{2} ||\mathbf{w}||^2}_{regularizer}$$

- Next on the menu:
 - Other regularlizers
 - Bayesian learning and quadratic regularizer

Other Regularizers



• Can use different norms for regularizer:

$$\tilde{E}(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} \{y(x_n, \mathbf{w}) - t_n\}^2 + \frac{\lambda}{2} \sum_{j=1}^{M} |w_j|^q$$

- e.g. q = 2 ridge regression
- e.g. q = 1 lasso
- · math is easiest with ridge regression

Optimization with a Quadratic Regularizer

• With q=2, total error still a nice quadratic:

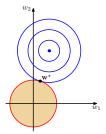
$$\tilde{E}(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} \{ y(x_n, \mathbf{w}) - t_n \}^2 + \frac{\lambda}{2} \mathbf{w}^T \mathbf{w}$$

Calculus ...

$$\mathbf{w} = (\underbrace{\lambda \mathbf{I} + \mathbf{\Phi}^T \mathbf{\Phi}}_{regularlized})^{-1} \mathbf{\Phi}^T \mathbf{t}$$

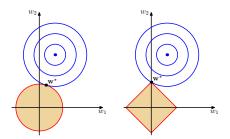
- Similar to unregularlized least squares
- Advantage $(\lambda \mathbf{I} + \mathbf{\Phi}^T \mathbf{\Phi})$ is well conditioned so inversion is stable

Ridge Regression vs. Lasso



- Ridge regression aka parameter shrinkage
 - Weights w shrink back towards origin

Ridge Regression vs. Lasso



- Ridge regression aka parameter shrinkage
 - Weights w shrink back towards origin
- Lasso leads to sparse models
 - Components of w tend to 0 with large λ (strong regularization)
 - Intuitively, once minimum achieved at large radius, minimum is on $w_1 = 0$

Outline

Regression

Linear Basis Function Models

Loss Functions for Regression

Finding Optimal Weights

Regularization

- Last week we saw an example of a Bayesian approach
 - Coin tossing prior on parameters
- We will now do the same for linear regression
 - Prior on parameter w
- There will turn out to be a connection to regularlization

- Start with a prior over parameters w
 - · Conjugate prior is a Gaussian:

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{0}, \alpha^{-1}\mathbf{I})$$

- This simple form will make math easier; can be done for arbitrary Gaussian too
- Data likelihood, Gaussian model as before:

$$p(t|\mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t|\mathbf{y}(\mathbf{x}, \mathbf{w}), \beta^{-1})$$

- Start with a prior over parameters w
 - Conjugate prior is a Gaussian:

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{0}, \alpha^{-1}\mathbf{I})$$

- This simple form will make math easier; can be done for arbitrary Gaussian too
- Data likelihood, Gaussian model as before:

$$p(t|\mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t|y(\mathbf{x}, \mathbf{w}), \beta^{-1})$$

• Posterior distribution on w:

$$p(\mathbf{w}|\mathbf{t}) \propto \left(\prod_{n=1}^{N} p(t_n|\mathbf{x}_n, \mathbf{w}, \beta)\right) p(\mathbf{w})$$

$$= \left[\prod_{n=1}^{N} \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp\left(-\frac{\beta}{2} (t_n - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n))^2\right) \right] \left(\frac{\alpha}{2\pi}\right)^{\frac{M}{2}} \exp\left(-\frac{\alpha}{2} \mathbf{w}^T \mathbf{w}\right)$$

Take the log:

$$-\ln p(\mathbf{w}|\mathbf{t}) = \frac{\beta}{2} \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2 + \frac{\alpha}{2} \mathbf{w}^T \mathbf{w} + const$$

 L₂ regularization is maximum a posteriori (MAP) with a Gaussian prior.

•
$$\lambda = \alpha/\beta$$

• Posterior distribution on w:

$$p(\mathbf{w}|\mathbf{t}) \propto \left(\prod_{n=1}^{N} p(t_n|\mathbf{x}_n, \mathbf{w}, \beta)\right) p(\mathbf{w})$$

$$= \left[\prod_{n=1}^{N} \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp\left(-\frac{\beta}{2} (t_n - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n))^2\right) \right] \left(\frac{\alpha}{2\pi}\right)^{\frac{M}{2}} \exp\left(-\frac{\alpha}{2} \mathbf{w}^T \mathbf{w}\right)$$

Take the log:

$$-\ln p(\mathbf{w}|\mathbf{t}) = \frac{\beta}{2} \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2 + \frac{\alpha}{2} \mathbf{w}^T \mathbf{w} + const$$

 L₂ regularization is maximum a posteriori (MAP) with a Gaussian prior.

•
$$\lambda = \alpha/\beta$$



Posterior distribution on w:

$$p(\mathbf{w}|\mathbf{t}) \propto \left(\prod_{n=1}^{N} p(t_n|\mathbf{x}_n, \mathbf{w}, \beta)\right) p(\mathbf{w})$$

$$= \left[\prod_{n=1}^{N} \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp\left(-\frac{\beta}{2} (t_n - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n))^2\right) \right] \left(\frac{\alpha}{2\pi}\right)^{\frac{M}{2}} \exp\left(-\frac{\alpha}{2} \mathbf{w}^T \mathbf{w}\right)$$

• Take the log:

$$-\ln p(\mathbf{w}|\mathbf{t}) = \frac{\beta}{2} \sum_{n=1}^{N} (t_n - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n))^2 + \frac{\alpha}{2} \mathbf{w}^T \mathbf{w} + const$$

 L₂ regularization is maximum a posteriori (MAP) with a Gaussian prior.

•
$$\lambda = \alpha/\beta$$



Posterior distribution on w:

$$p(\mathbf{w}|\mathbf{t}) \propto \left(\prod_{n=1}^{N} p(t_n|\mathbf{x}_n, \mathbf{w}, \beta)\right) p(\mathbf{w})$$

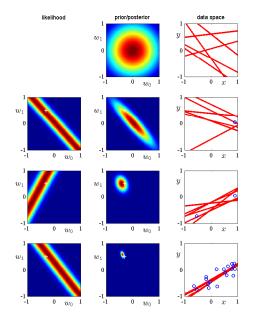
$$= \left[\prod_{n=1}^{N} \frac{\sqrt{\beta}}{\sqrt{2\pi}} \exp\left(-\frac{\beta}{2} (t_n - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n))^2\right) \right] \left(\frac{\alpha}{2\pi}\right)^{\frac{M}{2}} \exp\left(-\frac{\alpha}{2} \mathbf{w}^T \mathbf{w}\right)$$

Take the log:

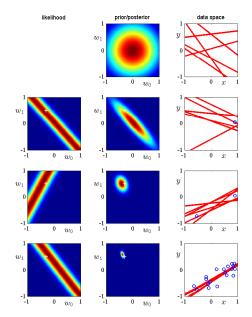
$$-\ln p(\mathbf{w}|\mathbf{t}) = \frac{\beta}{2} \sum_{n=1}^{N} (t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2 + \frac{\alpha}{2} \mathbf{w}^T \mathbf{w} + const$$

- L₂ regularization is maximum a posteriori (MAP) with a Gaussian prior.
 - $\lambda = \alpha/\beta$

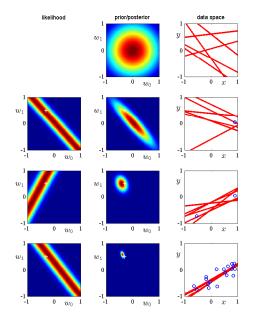




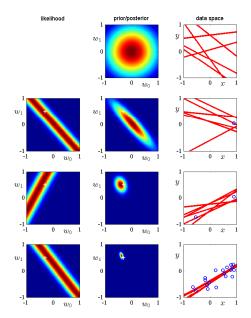
- Simple example $x, t \in \mathbb{R}$, $y(x, \mathbf{w}) = w_0 + w_1 x$
- parameter space
- Samples shown in data space
- Receive data points (blue circles in data space)
- Compute likelihood
- Posterior is prior (or prev. posterior) times likelihood



- Simple example $x, t \in \mathbb{R}$, $y(x, \mathbf{w}) = w_0 + w_1 x$
- Start with Gaussian prior in parameter space
- Samples shown in data space
- Receive data points (blue circles in data space)
- Compute likelihood
- Posterior is prior (or prev. posterior) times likelihood



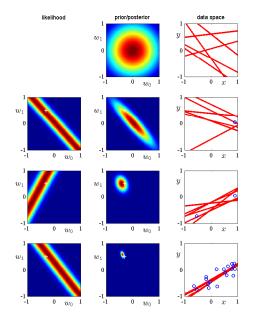
- Simple example $x, t \in \mathbb{R}$, $y(x, \mathbf{w}) = w_0 + w_1 x$
- Start with Gaussian prior in parameter space
- Samples shown in data space
- Receive data points (blue circles in data space)
- Compute likelihood
- Posterior is prior (or prev. posterior) times likelihood



- Simple example $x, t \in \mathbb{R}$, $y(x, \mathbf{w}) = w_0 + w_1 x$
- parameter space

Start with Gaussian prior in

- Samples shown in data space
- Receive data points (blue circles in data space)
- Compute likelihood
- Posterior is prior (or prev. posterior) times likelihood

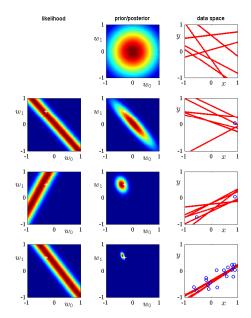


- Simple example $x, t \in \mathbb{R}$, $y(x, \mathbf{w}) = w_0 + w_1 x$
- parameter space

Start with Gaussian prior in

- Samples shown in data space
- Receive data points (blue circles in data space)
- Compute likelihood
- Posterior is prior (or prev. posterior) times likelihood

Bayesian Linear Regression - Intuition



- Simple example $x, t \in \mathbb{R}$, $y(x, \mathbf{w}) = w_0 + w_1 x$
- parameter space

Start with Gaussian prior in

- Samples shown in data space
- Receive data points (blue circles in data space)
- Compute likelihood
- Posterior is prior (or prev. posterior) times likelihood

- Single estimate of w (ML or MAP) doesn't tell whole story
- We have a distribution over w, and can use it to make predictions
- Given a new value for x, we can compute a distribution over t:

$$p(t|t,\alpha,\beta) = \int p(t,\mathbf{w}|t,\alpha,\beta)d\mathbf{w}$$

$$p(t|t,\alpha,\beta) = \int \underbrace{p(t|\mathbf{w},\beta)}_{predict} \underbrace{p(\mathbf{w}|t,\alpha,\beta)}_{probability} \underbrace{d\mathbf{w}}_{sum}$$

- i.e. For each value of w, let it make a prediction, multiply by its probability, sum over all w
- For arbitrary models as the distributions, this integral may not be computationally tractable

- Single estimate of w (ML or MAP) doesn't tell whole story
- We have a distribution over w, and can use it to make predictions
- Given a new value for x, we can compute a distribution over t:

$$p(t|\mathbf{t},\alpha,\beta) = \int p(t,\mathbf{w}|\mathbf{t},\alpha,\beta)d\mathbf{w}$$

$$p(t|t,\alpha,\beta) = \int \underbrace{p(t|w,\beta)}_{predict} \underbrace{p(w|t,\alpha,\beta)}_{probability} \underbrace{dw}_{sum}$$

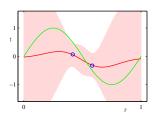
- i.e. For each value of w, let it make a prediction, multiply by its probability, sum over all w
- For arbitrary models as the distributions, this integral may not be computationally tractable

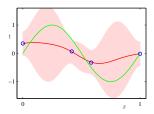
- Single estimate of w (ML or MAP) doesn't tell whole story
- We have a distribution over w, and can use it to make predictions
- Given a new value for x, we can compute a distribution over t:

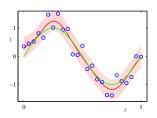
$$p(t|t,\alpha,\beta) = \int p(t,w|t,\alpha,\beta)dw$$

$$p(t|\mathbf{t},\alpha,\beta) = \int \underbrace{p(t|\mathbf{w},\beta)}_{predict} \underbrace{p(\mathbf{w}|\mathbf{t},\alpha,\beta)}_{probability} \underbrace{d\mathbf{w}}_{sum}$$

- i.e. For each value of w, let it make a prediction, multiply by its probability, sum over all w
- For arbitrary models as the distributions, this integral may not be computationally tractable







- With the Gaussians we've used for these distributions, the predicitve distribution will also be Gaussian
 - (math on convolutions of Gaussians)
- Green line is true (unobserved) curve, blue data points, red line is mean, pink one standard deviation
 - Uncertainty small around data points
 - · Pink region shrinks with more data

- So what do the Bayesians say about model selection?
 - Model selection is choosing model \mathcal{M}_i e.g. degree of polynomial, type of basis function ϕ
- Don't select, just integrate

$$p(t|\mathbf{x}, \mathcal{D}) = \sum_{i=1}^{L} \underbrace{p(t|\mathbf{x}, \mathcal{M}_i, \mathcal{D})}_{predictive \ dist.} p(\mathcal{M}_i|\mathcal{D})$$

- Average together the results of all models
- Could choose most likely model a posteriori $p(\mathcal{M}_i|\mathcal{D})$
 - More efficient, approximation

- So what do the Bayesians say about model selection?
 - Model selection is choosing model \mathcal{M}_i e.g. degree of polynomial, type of basis function ϕ
- Don't select, just integrate

$$p(t|\mathbf{x}, \mathcal{D}) = \sum_{i=1}^{L} \underbrace{p(t|\mathbf{x}, \mathcal{M}_i, \mathcal{D})}_{predictive \ dist.} p(\mathcal{M}_i|\mathcal{D})$$

- Average together the results of all models
- Could choose most likely model a posteriori $p(\mathcal{M}_i|\mathcal{D})$
 - More efficient, approximation

- So what do the Bayesians say about model selection?
 - Model selection is choosing model \mathcal{M}_i e.g. degree of polynomial, type of basis function ϕ
- Don't select, just integrate

$$p(t|\mathbf{x}, \mathcal{D}) = \sum_{i=1}^{L} \underbrace{p(t|\mathbf{x}, \mathcal{M}_i, \mathcal{D})}_{predictive \ dist.} p(\mathcal{M}_i|\mathcal{D})$$

- Average together the results of all models
- Could choose most likely model a posteriori $p(\mathcal{M}_i|\mathcal{D})$
 - More efficient, approximation

- So what do the Bayesians say about model selection?
 - Model selection is choosing model \mathcal{M}_i e.g. degree of polynomial, type of basis function ϕ
- Don't select, just integrate

$$p(t|\mathbf{x}, \mathcal{D}) = \sum_{i=1}^{L} \underbrace{p(t|\mathbf{x}, \mathcal{M}_i, \mathcal{D})}_{predictive \ dist.} p(\mathcal{M}_i|\mathcal{D})$$

- Average together the results of all models
- Could choose most likely model a posteriori $p(\mathcal{M}_i|\mathcal{D})$
 - · More efficient, approximation

• How do we compute the posterior over models?

$$p(\mathcal{M}_i|\mathcal{D}) \propto p(\mathcal{D}|\mathcal{M}_i)p(\mathcal{M}_i)$$

- Another likelihood + prior combination
- Likelihood:

$$p(\mathcal{D}|\mathcal{M}_i) = \int p(\mathcal{D}|\mathbf{w}, \mathcal{M}_i) p(\mathbf{w}|\mathcal{M}_i) d\mathbf{w}$$

How do we compute the posterior over models?

$$p(\mathcal{M}_i|\mathcal{D}) \propto p(\mathcal{D}|\mathcal{M}_i)p(\mathcal{M}_i)$$

- Another likelihood + prior combination
- Likelihood:

$$p(\mathcal{D}|\mathcal{M}_i) = \int p(\mathcal{D}|\mathbf{w}, \mathcal{M}_i) p(\mathbf{w}|\mathcal{M}_i) d\mathbf{w}$$

Conclusion

- Readings: Ch. 3.1, 3.1.1-3.1.4, 3.3.1-3.3.2, 3.4
- Linear Models for Regression
 - Linear combination of (non-linear) basis functions
- Fitting parameters of regression model
 - Least squares
 - Maximum likelihood (can be = least squares)
- Controlling over-fitting
 - Regularization
 - Bayesian, use prior (can be = regularization)
- Model selection
 - Cross-validation (use held-out data)
 - Bayesian (use model evidence, likelihood)